# Lecture 16: Estimation of Simultaneous Equations Models

Consider  $y_1 = Y_2 \gamma + X_1 \beta + \varepsilon_1$  which is an equation from a system.

We can rewrite this at  $y_1 = Z\delta + \varepsilon_1$  where  $Z = [Y_2 X_1]$ and  $\delta = [\gamma' \beta']'$ .

Note that  $Y_2$  is jointly determined with  $y_1$ , so

$$plim(1/N)Z'\varepsilon_1 \neq 0$$
 (usually).

### **IV Estimation**:

The point of IV estimation is to find a matrix of <u>instruments</u> W so that

$$\mathsf{plim}\frac{W'\varepsilon_1}{N} = \mathbf{0}$$

 $\quad \text{and} \quad$ 

$$\mathsf{plim}\frac{W'Z}{N} = Q$$

where Q is nonsingular.

The IV estimator is  $(W'Z)^{-1}W'y_1$ . As in the lecture on dynamic models, multiplying the model by the transpose of the matrix of instruments yields  $W'y_1 = W'Z\delta +$  $W'\varepsilon_1$  which gives  $\hat{\delta}_{IV}$ .

Asymptotic distribution of  $\hat{\delta}_{IV}$ :

Note that 
$$\hat{\delta}_{IV} - \delta = (W'Z)^{-1}W'\varepsilon_1$$
. Assume that  
 $\frac{W'\varepsilon_1}{\sqrt{N}} \to N\left(0, \sigma^2 \frac{W'W}{N}\right).$ 

(Is this a sensible assumption? Recall the CLT.)

Then

$$\sqrt{N}(\hat{\delta}_{IV} - \delta) \rightarrow N(\mathbf{0}, \sigma^2 \sum_{\delta})$$

#### where

$$\sum_{\delta} = N(W'Z)^{-1}W'W(W'Z)^{-1} = (1/N)Q^{-1}W'WQ^{-1}$$

The question is what to use for W. Suppose we use X.

Multiplying by the transpose of the matrix of instruments gives  $X'y_1 = X'Z\delta + X'\varepsilon_1$ .

For this system of equations to have a solution, X'Z has to be square and nonsingular. When is this possible?

Note the following dimensions: X is  $N \times K$ ,  $X_1$  is  $N \times K_1$  and  $Y_2$  is  $N \times (G_1 - 1)$ . This, of course, requires  $K = K_1 + G_1 - 1$ .

(*Recall the* <u>order condition</u>:  $K \ge K_1 + G_1 - 1$ .)

The resulting IV estimates are indirect least squares which we saw last time.

Suppose  $K < K_1 + G_1 - 1$ . Then what happens? Consider the supply and demand example. This is the underidentified case.

Suppose  $K > K_1 + G_1 - 1$ . Then  $X'y_1 = X'Z\delta + X'\varepsilon_1$  is K equations in  $K_1 + G_1 - 1$  unknowns (setting  $X'\varepsilon_1$  to 0 which is its expectation). We could choose  $K_1 + G_1 - 1$  equations to solve for  $\delta$ - there are many ways to do this, typically leading to different estimates. This is the overidentified case.

Another way to look at this case is as a regression model - with K "observations" on the dependent variable.

We could apply the LS method, but the GLS is more efficient since  $V(X'\varepsilon_1) = \sigma^2(X'X) (\neq \sigma^2 I)$ .

The observation matrix is  $X'y_1$  and X'Z. GLS gives the estimator

$$\hat{\delta} = [Z'X(X'X)^{-1}X'Z]^{-1}Z'X(X'X)^{-1}X'y_1.$$

In the just-identified case (where X'Z is invertible),

$$\hat{\delta} = (X'Z)^{-1}X'X(Z'X)^{-1}Z'X(X'X)^{-1}X'y_1 = (X'Z)^{-1}X'y_1 = \hat{\delta}_{IV} \text{ with } W = X.$$

## TWO-STAGE LEAST SQUARES:

Return to the <u>overidentified</u> case:

$$\hat{\delta} = [Z'X(X'X)^{-1}X'Z]^{-1}Z'X(X'X)^{-1}X'y_1.$$

*Proposition*: The estimator

$$\hat{\delta} = [Z'X(X'X)^{-1}X'Z]^{-1}Z'X(X'X)^{-1}X'y_1$$

is the two-stage least squares (2SLS or TSLS) estimator.

Why is  $\hat{\delta}$  called the TSLS estimator?

Let 
$$\overline{M} = X(X'X)^{-1}X' = I - M$$
.

Then 
$$\hat{\delta} = (Z'ar{M}Z)^{-1}Z'ar{M}y_1$$
 .

We will write out the expression for  $\hat{\delta}$ .

$$\hat{\delta} = \begin{bmatrix} \hat{Y}_{2}'\hat{Y}_{2} & \hat{Y}_{2}'X_{1} \\ X_{1}'\hat{Y}_{2} & X_{1}'X_{1} \end{bmatrix}^{-1} \begin{bmatrix} \hat{Y}_{2}'y_{1} \\ X_{1}'y_{1} \end{bmatrix}.$$

Now:  $\overline{M}Y_2 = X(X'X)^{-1}X'Y_2 = \hat{Y}_2 = X\hat{\Pi}_2$  which is the LS predictor of  $Y_2$ .

$$Z'\bar{M}Z = \begin{bmatrix} Y_2'\bar{M}Y_2 & Y_2'\bar{M}X_1\\ X_1'\bar{M}Y_2 & X_1'\bar{M}X_1 \end{bmatrix}$$

Note that  $X'_1 \overline{M} X_1 = X'_1 X_1 \cdot \left( R[X_1] \subset R[X] \Rightarrow \overline{M} X_1 = X_1; R \in \mathbb{N} \right)$ 

Also: 
$$Y'_2 \overline{M} Y_2 = Y'_2 \overline{M} \overline{M} Y_2 = \hat{Y}'_2 \hat{Y}_2.$$

So,

$$\hat{\delta} = \begin{bmatrix} \hat{Y}_2'\hat{Y}_2 & \hat{Y}_2'X_1\\ X_1'\hat{Y}_2 & X_1'X_1 \end{bmatrix}^{-1} \begin{bmatrix} \hat{Y}_2'y_1\\ X_1'y_1 \end{bmatrix}$$

 $\hat{\delta}$  is the coefficient vector from a regression of  $y_1$  on  $\hat{Y}_2$  and  $X_1$ .

Interpretation as 2SLS? Interpretation as IV?

*Proposition*: 2SLS is IV estimation with  $W = [\hat{Y}_2 X_1]$ .

Proof: Note that  $W'Z = \begin{bmatrix} \hat{Y}_2'Y_2 & \hat{Y}_2'X_1 \\ X_1'\hat{Y}_2 & X_1'X_1 \end{bmatrix} = \begin{bmatrix} \hat{Y}_2'\hat{Y}_2 & \hat{Y}_2'X_1 \\ X_1'\hat{Y}_2 & X_1'X_1 \end{bmatrix}.$ 

This is the matrix appearing inverted in  $\hat{\delta}$ .

Asymptotic distribution of  $\hat{\delta}$ : We know this from IV results.

Note that  $\hat{\delta} = \delta + (Z'\bar{M}Z)^{-1}Z'\bar{M}\varepsilon_1$ . The asymptotic variance of  $N^{1/2}(\hat{\delta} - \delta)$  is the asymptotic variance of  $N^{1/2}(Z'\bar{M}Z)^{-1}Z'\bar{M}\varepsilon_1 = u$ .

$$Var(u) = N\sigma^{2}(Z'\bar{M}Z)^{-1}Z'\bar{M}'Z(Z'\bar{M}Z)^{-1} = N\sigma^{2}(Z'\bar{M}Z)^{-1}.$$

Remember to remove the N in calculating estimated variance for  $\hat{\delta}$ . (Why?)

Estimation of  $\sigma^2$ :

$$\hat{\sigma}^2 = (y_1 - Z\hat{\delta})'(y_1 - Z\hat{\delta})/N.$$

Note that  $Z = [Y_2X_1]$  appears in the expressions for  $\hat{\sigma}^2$ , <u>not</u>  $[\hat{Y}_2X_1]$ .

If you regress  $y_1$  on  $\hat{Y}_2$  and  $X_1$ , you will get the right coefficients but the wrong standard errors.

## **GEOMETRY OF 2SLS**:

Take:

- N = 3 (observations)
- K = 2 (exogenous variables),
- $K_1 = 1$  (included exogenous variables) and

 $G_1 = 2$  (included endogenous variables - one is normalized).

How many parameters?



 $\hat{Y}_2$  is in the plane spanned by  $X_1$  and  $X_2$ .  $y_1$  is projected to the plane spanned by  $\hat{Y}_2$  and  $X_1$ .

Note that  $X_1$  and  $X_2$  and  $X_1$  and  $\hat{Y}_2$  span the same plane. (*Why*?)

Model is just identified (projection of both stages is to the same plane).

What happens if the model is <u>overidentified</u>? (For example,  $K_1 = 0$ , that is, no included regressors).

What if <u>underidentified</u>? (For example,  $K_2 = 2$ , that is, no excluded regressors).