# Economics 620, Lecture 8: Asymptotics I

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We are interested in the properties of estimators as

 $n \to \infty$ .

Consider a sequence of random variables

$$\{X_n, n\geq 1\}.$$

Often  $X_n$  is an estimator such as a sample mean or  $\widehat{\beta}_n$ 

Often it is convenient to center the sequence:  $\{\widehat{\beta}_n - \beta\}$ 

and sometimes to scale  $\{(\widehat{\beta_n} - \beta)/\sigma_n\}$ 

### Definition: (convergence in probability)

A sequence of random variables  $\{X_n, n \ge 1\}$  is said to *converge weakly* to a constant *c* if

$$\lim_{n\to\infty}P(|X_n-c|>\varepsilon)=0$$

for every given  $\varepsilon > 0$ .

This is written  $p \lim X_n = c$  or  $X_n \xrightarrow{p} c$ 

Some properties of plim:

1. plim XY =plim X plim Y

2. plim 
$$(X + Y) = plim X + plim Y$$

3. Slutsky's theorem: If the function g is continuous at plim X, then plim g(X) = g(plim X).

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### Definition: (Strong convergence)

A sequence of random variables is said to *converge strongly* to a constant c if

$$P(\lim_{n\to\infty}X_n=c)=1$$

or

$$\lim_{N\to\infty} P(\sup_{n>N} |x_n-c| > \varepsilon) = 0.$$

Strong convergence is also called *almost sure convergence* or *convergence* with probability one and is written  $X_n \to c$  w.p. 1 or  $X_n \stackrel{a.s.}{\to} c$ .

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# Difference betwen convergence a.s. and plim

plim involves probabilities on each element of the sequence, and limits of these probabilities.

$$\lim_{n\to\infty} P(|X_n-c|>\varepsilon)=0$$

Strong convergence involves probabilities on the entire sequence.

$$P(\lim_{n\to\infty}X_n=c)=1$$

Sequence of marginal probabilities vs. joint probability over infinite sequences.

Note a.s. convergence implies plim.

Difference usually doesn't matter in applications and plim is easier to establish.

Let  $\{X_n, n \ge 1\}$  be observations and suppose we look at the sequence

$$\bar{X}_n = \sum_{i=1}^n X_i / n$$

when does  $\bar{X}_n \xrightarrow{p} \xi$  where  $\xi$  is some parameter?

Weak Law of Large Numbers: (WLLN) Let  $E(X_i) = \mu$ ,  $V(X_i) = \sigma^2$ ,  $cov(X_iX_j) = 0$ . Then  $\bar{X}_n - \mu \to 0$  in probability.

Lemma: Chebyshev's Inequality:

 $P(|X - \mu| \ge k) \le \sigma^2/k^2$  where  $E(X) = \mu$  and  $V(X) = \sigma^2$ .

Proof of Chebshev's inequality

$$\sigma^2 = \int (x - \mu)^2 dF$$

$$=\int_{-\infty}^{\mu-k} (x-\mu)^2 dF + \int_{\mu-k}^{\mu+k} (x-\mu)^2 dF$$
$$+ \int_{\mu+k}^{\infty} (x-\mu)^2 dF$$

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Put in the largest value of x in the first and smallest in the last integral, and drop the middle to get:

$$\sigma^2 \ge k^2 P(|x - \mu| \ge k)$$

Proof of WLLN: Since we are interested in  $\overline{X_n}$ , note that

$$E(\overline{X_n}) = \mu$$
 and  $V(\overline{X_n}) = \sigma^2/n$ .

Consequently,

$$\lim_{n\to\infty} P(|\bar{X}_n - \mu| > \varepsilon) \leq \lim_{n\to\infty} \sigma^2 / n\varepsilon^2 = 0.$$

## Notes:

1.  $E(X_i) = \mu_i$  is okay. Consider

$$ar{X}_n - ar{\mu}_n$$
 with  $ar{\mu}_n = n^{-1} \sum \mu_i$ .

2.  $V(X_i) = \sigma_i^2$  is okay. As long as  $\lim \sum \sigma_i^2/n^2 = 0$ , our proof applies.

3. Existence of  $\sigma^2$  can be dropped if we assume independent and identically distributed observations.

In this case, the proof is different and is based on Markov's inequality

$$P(|X| \ge k) \le E|X|/k$$

from which Chebyshev's inequality follows.

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# Strong Law of Large Numbers:

If  $X_i$  are *independent* with  $E(X_i) = \mu_i$ ,  $V(X_i) = \sigma_i^2$  and  $\sum \sigma_i^2/i^2 < \infty$ . Then  $\bar{X}_n - \bar{\mu}_n \to 0$  almost surely (a.s.).

We can drop the existence of  $\sigma_i^2$  if we assume independent and identically distributed observations.

Example (Shiryayev): let the probability space be [0,1) with Lebesgue measure (length of intervals). To each element  $\omega$  of [0,1), there is a sequence  $\{x_i\}$  where  $x_i$  is the ith element in the dyadic expansion of  $\omega$ , i.e.  $\omega=0.x_1x_2..., x_i \in \{0,1\}$ . Then  $P(X_1 = x_1, ..., X_n = x_n) = P(x_1/2 + x_2/2^2 + ...x_n/2^n \le \omega < x_1/2 + x_2/2^2 + ...x_n/2^n + 1/2^n) = 1/2^n$ . Thus P(X = 1) = P(X = 0) = 1/2 and the obs are iid. By the SLLN,  $\sum_{i=1}^{n} X_i/n \to 1/2$ .

Interpretation? Borel result on normal numbers.

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#### Definition: (Convergence in distribution):

A sequence of random variables  $\{X_n, n \ge 1\}$  with distribution functions  $\{F_n(x) = P(X_n \le x), n \ge 1\}$  is said to *converge in distribution* to a random variable X with distribution function F(x) if and only if  $\lim_{n\to\infty} F_n(x) = F(x)$  at all points of continuity of F(x).

Notation:  $X_n \xrightarrow{D} X$ . plim is a special case in which F is a degenerate distribution.

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An equivalent characterization is:

$$Ef(X_n) \to Ef(X)$$

for all bounded continuous functions f.Another is

$$P(X_n \in B) \rightarrow P(X \in B)$$

for all sets B with  $P(\partial B) = 0$ .

We have 
$$X_n \xrightarrow{as} X_n \xrightarrow{p} X_n \xrightarrow{d}$$

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Convergence in distribution is used to approximate the distribution of estimators.

If an estimator is consistent (plim=true value), studying the limiting distribution nontrivially requires norming.

**CMT**: Let g(x) be continuous on a set which has probability one. Then

$$X_{n} \stackrel{d}{\to} X \Rightarrow g(X_{n}) \stackrel{d}{\to} g(X)$$
$$X_{n} \stackrel{p}{\to} X \Rightarrow g(X_{n}) \stackrel{p}{\to} g(X)$$
$$X_{n} \stackrel{as}{\to} X \Rightarrow g(X_{n}) \stackrel{as}{\to} g(X)$$

The CMT is extremely useful. Why?

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# Some properties of convergence in probability (plim) and convergence in distribution:

1.  $X_n$  and  $Y_n$  are random variable sequences. If  $plim(X_n - Y_n) = 0$ and  $Y_n \xrightarrow{D} Y$ , then  $X_n \xrightarrow{D} Y$  as well. This is an extremely useful device.

2. If 
$$Y_n \xrightarrow{D} Y$$
 and  $X_n \to c$  in probability (i.e., plim $X_n = c$ ), then

a. 
$$X_n + Y_n \xrightarrow{D} c + Y$$

b. 
$$X_n Y_n \xrightarrow{D} cY$$

c. 
$$Y_n/X_n \xrightarrow{D} Y/c, c \neq 0.$$

This notation is used to denote relative orders of magnitude of sequences in the limit.

Sequences  $\{x_i\}, \{b_i\}$  (nonstochastic, for now)

$$x_n = O(b_n) \Rightarrow \lim_{n \to \infty} x_n/b_n = -\infty < c < \infty$$
  
 $x_n = o(b_n) \Rightarrow \lim_{n \to \infty} x_n/b_n = 0$ 

Thus

$$egin{aligned} & x_n = o(1) \Rightarrow x_n 
ightarrow 0; x_n = o(n) \Rightarrow x_n/n 
ightarrow 0 \ & x_n = O(1) \Rightarrow x_n 
ightarrow c; x_n = O(n) \Rightarrow x_n/n 
ightarrow c \end{aligned}$$

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For stochastic sequences  $\{X_i\}$  we have

$$X_n = O_p(b_n) \Rightarrow orall \epsilon \exists C$$
 such that  $\lim_{n o \infty} P(|X_n/b_n| < C) > 1 - \epsilon$ 

This says that the ratio remains bounded in probability. Also

$$X_n = o_p(b_n) \Rightarrow p \lim X_n/b_n = 0$$

Thus for example (using results above) if

$$X_n - Y_n = o_p(1)$$
 and  $X_n \stackrel{d}{\rightarrow} X$ 

Then  $Y_n \xrightarrow{d} X$ 

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# Further Properties of $O_p$ and $o_p$ $o_p(1) + o_p(1) = o_p(1)$ $o_p(1) + O_p(1) = O_p(1)$ $O_p(1)o_p(1) = o_p(1)$ $(1 + o_p(1))^{-1} = O_p(1)$ $o_p(b_n) = b_n o_p(1)$ $O_p(b_n) = b_n O_p(1)$

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