Economics 620, Lecture 16: Estimation of Simultaneous Equations Models

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Consider $y_1 = Y_2 \gamma + X_1 \beta + \varepsilon_1$ which is an equation from a system.

We can rewrite this at $y_1 = Z\delta + \varepsilon_1$ where $Z = [Y_2 \ X_1]$ and $\delta = [\gamma' \ \beta']'$.

Note that Y_2 is jointly determined with y_1 , so

$$plim(1/N)Z'\varepsilon_1 \neq 0$$
 (usually).

IV Estimation:

The point of IV estimation is to find a matrix of instruments W so that

$$plim \frac{W'\varepsilon_1}{N} = 0$$

and

$$plim \frac{W'Z}{N} = Q$$

where Q is nonsingular.

The IV estimator is $(W'Z)^{-1}W'y_1$. As in the lecture on dynamic models, multiplying the model by the transpose of the matrix of instruments yields $W'y_1 = W'Z\delta + W'\varepsilon_1$ which gives $\hat{\delta}_{IV}$.

Asymptotic distribution of $\hat{\delta}_{IV}$:

Note that $\hat{\delta}_{IV} - \delta = (W'Z)^{-1}W'\varepsilon_1$. Assume that

$$\frac{W'\varepsilon_1}{\sqrt{N}} \to N\left(0, \sigma^2 \frac{W'W}{N}\right).$$

(Is this a sensible assumption? Recall the CLT.)

Then

$$\sqrt{N}(\hat{\delta}_{IV} - \delta) \rightarrow N(0, \sigma^2 \sum_{\delta})$$

where

$$\sum_{\delta} = N(W'Z)^{-1}W'W(W'Z)^{-1} = (1/N)Q^{-1}W'WQ^{-1}.$$

The question is what to use for W. Suppose we use X.

Multiplying by the transpose of the matrix of instruments gives $X'y_1 = X'Z\delta + X'\varepsilon_1$.

For this system of equations to have a solution, X'Z has to be square and nonsingular. When is this possible?

Note the following dimensions: X is $N \times K$, X_1 is $N \times K_1$ and Y_2 is $N \times (G_1 - 1)$. This, of course, requires $K = K_1 + G_1 - 1$. (Recall the order condition: $K \ge K_1 + G_1 - 1$.)

Thus, the above procedure works when the equation is just identified.

The resulting IV estimates are indirect least squares which we saw last time.

Suppose $K < K_1 + G_1 - 1$. Then what happens? Consider the supply and demand example. This is the underidentified case.

Suppose $K > K_1 + G_1 - 1$. Then $X'y_1 = X'Z\delta + X'\varepsilon_1$ is K equations in $K_1 + G_1 - 1$ unknowns (setting $X'\varepsilon_1$ to 0 which is its expectation). We could choose $K_1 + G_1 - 1$ equations to solve for δ - there are many ways to do this, typically leading to different estimates. This is the overidentified case.

Another way to look at this case is as a regression model - with K "observations" on the dependent variable.

We could apply the LS method, but the GLS is more efficient since $V(X'\varepsilon_1) = \sigma^2(X'X)(\neq \sigma^2I)$.

The observation matrix is $X'y_1$ and X'Z. GLS gives the estimator

$$\hat{\delta} = [Z'X(X'X)^{-1}X'Z]^{-1}Z'X(X'X)^{-1}X'y_1.$$

In the just-identified case (where X'Z is invertible),

$$\hat{\delta} = (X'Z)^{-1}X'X(Z'X)^{-1}Z'X(X'X)^{-1}X'y_1$$

= $(X'Z)^{-1}X'y_1 = \hat{\delta}_{IV}$ with $W = X$.

TWO-STAGE LEAST SQUARES:

Return to the overidentified case:

$$\hat{\delta} = [Z'X(X'X)^{-1}X'Z]^{-1}Z'X(X'X)^{-1}X'y_1.$$

Proposition: The estimator

$$\hat{\delta} = [Z'X(X'X)^{-1}X'Z]^{-1}Z'X(X'X)^{-1}X'y_1$$

is the two-stage least squares (2SLS or TSLS) estimator.

Why is $\hat{\delta}$ called the TSLS estimator?

Let
$$\bar{M} = X(X'X)^{-1}X' = I - M$$
.
Then $\hat{\delta} = (Z'\bar{M}Z)^{-1}Z'\bar{M}y_1$.

We will write out the expression for $\hat{\delta}$.

$$\hat{\delta} = \begin{bmatrix} \hat{Y}_2' \hat{Y}_2 & \hat{Y}_2' X_1 \\ X_1' \hat{Y}_2 & X_1' X_1 \end{bmatrix}^{-1} \begin{bmatrix} \hat{Y}_2' y_1 \\ X_1' y_1 \end{bmatrix}.$$

Now: $\overline{M}Y_2 = X(X'X)^{-1}X'Y_2 = \hat{Y}_2 = X\hat{\Pi}_2$ which is the LS predictor of Y_2 .

$$Z'\bar{M}Z = \left[\begin{array}{cc} Y_2'\bar{M}Y_2 & Y_2'\bar{M}X_1 \\ X_1'\bar{M}Y_2 & X_1'\bar{M}X_1 \end{array} \right]$$

Note that $X_1'\bar{M}X_1=X_1'X_1.\big(R[X_1]\subset R[X]\Rightarrow \bar{M}X_1=X_1;\bar{M}X=X\big).$

Also: $Y_2' \bar{M} Y_2 = Y_2' \bar{M} \bar{M} Y_2 = \hat{Y}_2' \hat{Y}_2$.

So,

$$\hat{\delta} = \begin{bmatrix} \hat{Y}_2' \hat{Y}_2 & \hat{Y}_2' X_1 \\ X_1' \hat{Y}_2 & X_1' X_1 \end{bmatrix}^{-1} \begin{bmatrix} \hat{Y}_2' y_1 \\ X_1' y_1 \end{bmatrix}.$$

 $\hat{\delta}$ is the coefficient vector from a regression of y_1 on \hat{Y}_2 and X_1 .

Interpretation as 2SLS? Interpretation as IV?

Proposition: 2SLS is IV estimation with $W = [\hat{Y}_2 X_1]$.

Proof: Note that

$$W'Z = \begin{bmatrix} \hat{Y}_2'Y_2 & \hat{Y}_2'X_1 \\ X_1'\hat{Y}_2 & X_1'X_1 \end{bmatrix} = \begin{bmatrix} \hat{Y}_2'\hat{Y}_2 & \hat{Y}_2'X_1 \\ X_1'\hat{Y}_2 & X_1'X_1 \end{bmatrix}.$$

This is the matrix appearing inverted in $\hat{\delta}$.

Asymptotic distribution of $\hat{\delta}$: We know this from IV results.

Note that $\hat{\delta} = \delta + (Z'\bar{M}Z)^{-1}Z'\bar{M}\varepsilon_1$. The asymptotic variance of $N^{1/2}(\hat{\delta} - \delta)$ is the asymptotic variance of $N^{1/2}(Z'\bar{M}Z)^{-1}Z'\bar{M}\varepsilon_1 = u$.

$$Var(u) = N\sigma^2 (Z'\bar{M}Z)^{-1} Z'\bar{M}Z (Z'\bar{M}Z)^{-1}]$$

= $N\sigma^2 (Z'\bar{M}Z)^{-1}$.

Remember to remove the N in calculating estimated variance for $\hat{\delta}$. (Why?)

Estimation of σ^2 :

$$\hat{\sigma}^2 = (y_1 - Z\hat{\delta})'(y_1 - Z\hat{\delta})/N.$$

Note that $Z = [Y_2X_1]$ appears in the expressions for $\hat{\sigma}^2$, <u>not</u> $[\hat{Y}_2X_1]$.

If you regress y_1 on \hat{Y}_2 and X_1 , you will get the right coefficients but the wrong standard errors.

GEOMETRY OF 2SLS:

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Take:
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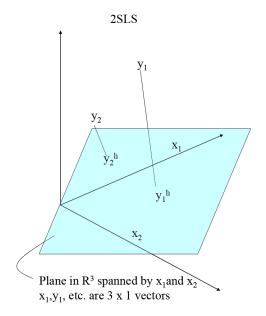
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N = 3 (observations)
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K = 2 (exogenous variables),

 $K_1 = 1$ (included exogenous variables) and

 $G_1 = 2$ (included endogenous variables - one is normalized).

How many parameters?



 \hat{Y}_2 is in the plane spanned by X_1 and X_2 . y_1 is projected to the plane spanned by \hat{Y}_2 and X_1 .

Note that X_1 and X_2 and X_1 and \hat{Y}_2 span the same plane. (Why?)

Model is just identified (projection of both stages is to the same plane).

What happens if the model is <u>overidentified</u>? (For example, $K_1=0$, that is, no included regressors).

What if <u>underidentified</u>? (For example, $K_2 = 2$, that is, no excluded regressors).